



Cardiovascular Risk in Users of Mirabegron Compared with Users of Antimuscarinic Treatments for Overactive Bladder: Findings from a Non-Interventional, Multinational, Cohort Study

Veena Hoffman¹ · Jesper Hallas² · Marie Linder³ · Andrea V. Margulis⁴ · Brandon T. Suehs⁵ · Alejandro Arana⁴ · Kelesitse Phiri¹ · Cheryl Enger¹ · Libby Horter⁵ · Ingvild Odsbu³ · Morten Olesen² · Susana Perez-Gutthann⁴ · Yihua Xu⁵ · Nina Sahlertz Kristiansen² · Kwame Appenteng⁶ · Stefan de Vogel⁷ · John D. Seeger¹ on behalf of the mirabegron PMR-PASS study group

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Abstract

Introduction During clinical trials, mirabegron, a β_3 -adrenoreceptor agonist, was associated with increased vital signs vs placebo in patients with overactive bladder.

Objective The purpose of this study was to compare incidence rates of adverse cardiovascular (CV) outcomes following mirabegron or antimuscarinic use.

Methods We conducted an observational post-marketing safety study utilising real-world data. The study population was identified within five sources: Danish and Swedish National Registers, Clinical Practice Research Datalink (UK), Optum (USA) and Humana (USA). Episodes of time when patients were new users of mirabegron or antimuscarinics (October 2012–December 2018) were sourced from prescriptions and matched on propensity scores. Occurrences of major adverse cardiovascular events (MACE), acute myocardial infarction (AMI), stroke, CV mortality and all-cause mortality were identified. Outcome incidence rates and hazard ratios from Cox models were estimated.

Results Overall, 152,026 mirabegron and 152,026 antimuscarinic episodes were matched. The population consisted of 63.1% women and 72.6% were ≥ 65 years old. There were no appreciable differences in the incidence rates of MACE, AMI or stroke between users of mirabegron and antimuscarinics. Incidence rates of CV mortality (hazard ratio 0.83, 95% confidence interval 0.73–0.95) and all-cause mortality (hazard ratio 0.80, 95% confidence interval 0.76–0.84) were no higher with mirabegron vs antimuscarinics. Results restricted to episodes at high risk for CV events or stratified by age (< 65 years, ≥ 65 years) or prior overactive bladder medication use were consistent with overall findings.

Conclusions This large, multinational study found no higher risk of MACE, AMI, stroke, CV mortality or all-cause mortality among users of mirabegron relative to users of antimuscarinics.

Plain Language Summary

During clinical trials, mirabegron, which is a treatment for overactive bladder, was associated with small increases in heart rate and blood pressure. This study was conducted to compare the frequency of cardiac events following the use of mirabegron or antimuscarinics, a group of treatments also used to treat overactive bladder. We obtained the data for this study from four countries: Denmark, Sweden, the UK and the USA. We identified people who were new users of mirabegron or antimuscarinics from 2012 to 2018 using prescription or dispensing records. Occurrences of major cardiac events, heart attack, stroke, death due to cardiac events and death from any cause were evaluated. Overall, we identified 152,026 times when mirabegron or antimuscarinics were each used as new treatments. Most of the people in the study were women (63.1%) and

The mirabegron PMR-PASS study group members are listed in the Acknowledgements section.

✉ John D. Seeger
John.Seeger@optum.com

Extended author information available on the last page of the article

at least 65 years old (72.6%). There were no notable differences between the treatment groups with regard to how often major cardiac events, heart attack or stroke occurred. Further, death due to cardiac events and death from any cause were no higher with mirabegron compared with antimuscarinics. We obtained similar results when the data were assessed for patients who were at high risk for cardiac events or split by age (less than 65 years or at least 65 years) or a history of overactive bladder medication use. In conclusion, this large study involving data from several countries found no higher risk of major cardiac events, heart attack, stroke or death among people prescribed mirabegron compared with antimuscarinics.

Key Points

This study was conducted to compare the incidence of adverse cardiovascular (CV) outcomes following the use of two types of treatment for overactive bladder, mirabegron and antimuscarinics.

In total, 152,026 occurrences of new use for each of mirabegron or antimuscarinics were matched. No higher risk of major adverse CV events, acute myocardial infarction, stroke, CV mortality or all-cause mortality was noted among current users of mirabegron relative to current users of antimuscarinics.

Similar results to the overall findings were observed when the data were restricted to patients who were at high risk for CV events or stratified by age (< 65 years, ≥ 65 years) or prior overactive bladder medication use.

1 Introduction

Overactive bladder (OAB) syndrome is a common condition with an estimated prevalence between 11 and 36% [1–3]. The syndrome is defined as the presence of urinary urgency, with or without urinary incontinence, generally with increased daytime frequency and nocturia, without any proven infection or other obvious pathology [4]. The effect of OAB on the quality of life of affected individuals is well understood, with patients reporting increased levels of anxiety and depression, as well as impairments in sexual activity and their ability to work [5, 6].

Several phase III clinical trials investigated the safety and efficacy of the β_3 -adrenoreceptor agonist, mirabegron, for treating patients with OAB [7–9]. The results from these trials informed the approval of mirabegron for the treatment of patients with OAB symptoms by the US Food and Drug Administration and the European Medicines Agency in 2012 [10, 11]. Findings obtained during the OAB clinical programme indicated that mirabegron, administered at a once daily dose of 50 mg, was associated with an increase in pulse rate of approximately 1 beat per minute and a

systolic blood pressure/diastolic blood pressure increase of 0.4 to 0.6 mmHg compared with placebo [12]. An electrocardiogram study also found that a supratherapeutic dose of mirabegron 200 mg caused QT prolongation in female subjects [13]. In the context of the vital sign observations, regulatory authorities requested that a post-authorisation safety study (PASS) be conducted to address concerns about the potential cardiovascular (CV) safety of mirabegron. The current study was conducted to address this request using antimuscarinics as an active comparator group.

2 Methods

2.1 Data Sources

This study used five data sources from four countries and was a collaborative effort by partners from the University of Southern Denmark (Danish National Registers); the Centre for Pharmacoepidemiology, Karolinska Institutet (Swedish National Registers); RTI Health Solutions (Clinical Practice Research Datalink [CPRD]; UK); Optum (Optum Research Database [ORD]; USA) and Humana Healthcare Research (Humana Database; USA). The CPRD population was analysed as two groups, CPRD-linked (data from general practices that permitted hospital and mortality data linkage) and CPRD-unlinked (no linkage). The two US data sources were linked to the National Death Index.

Each research partner followed the same core protocol and statistical analysis plan, although methodological details varied at each site to account for differing data environments (see Electronic Supplementary Material [ESM]). Within each site, standard operating procedures and quality-control measures guided study conduct.

2.2 Study Design

This non-interventional, cohort study using secondary data included patients exposed to mirabegron or antimuscarinic medications from October 2012 through December 2018. Study periods within each source varied according to data availability. Exposure was based on prescription or dispensing data (hereafter prescriptions). The analysis incorporated a new user design and was conducted at the episode-of-use level as switching was expected to occur. Each mirabegron

episode was matched to one antimuscarinic episode that started in the same year within each data source and according to year-specific propensity scores (PSs), with forced matching on age and sex. Incidence rates of major adverse cardiovascular events (MACE), acute myocardial infarction (AMI), stroke, CV mortality and all-cause mortality were estimated within person-time of current exposure of the matched mirabegron and antimuscarinic episodes. Hazard ratios (HRs) were obtained from Cox proportional hazard models that compared mirabegron with antimuscarinics.

2.3 Inclusion and Exclusion Criteria

Episode inclusion criteria were: a prescription for mirabegron or antimuscarinics (darifenacin, fesoterodine, oxybutynin, solifenacin, tolterodine, trospium) during the study period with no previous use of that specific medication in the prior 12 months (new use). Trospium was not available for use in Sweden during the study period. Additionally, patients were required to be ≥ 18 years old and have ≥ 12 months of prior enrolment in the data source. Prescriptions of two study medications on the same day and non-tablet antimuscarinics were excluded. Patients could contribute multiple episodes if each episode met the above criteria. Additional site-specific inclusion and exclusion criteria were implemented without impacting the overall study design (ESM).

2.4 Exposure Definition

Each person-day was classified as currently exposed to medication if the person-day was within the days of supply of the prescription plus a grace period of 50% [14]. The grace period accounted for variable adherence to medications where a patient still had some medication available beyond the days of supply due to forgetting some scheduled doses or changes in dosing schedule. Methods for calculating total days of supply varied according to the information available from each data source (ESM).

Current person-time for a treatment group was terminated upon a prescription of a medication from the other treatment group or following the end of the days of supply (after applying the grace period). When days of supply overlapped between prescriptions, the first prescription was truncated on the day before the subsequent prescription. Any patient could contribute current exposure person-time for mirabegron and antimuscarinics if they switched during the study period, but not simultaneously.

2.5 Study Outcomes

The main outcome was MACE, defined as the first occurrence of AMI, stroke or CV mortality. Additional outcomes included the individual events of AMI, stroke, CV mortality

and all-cause mortality. AMI outcomes included both ST elevation myocardial infarction and non-ST elevation myocardial infarction [15] and stroke outcomes included haemorrhagic and ischaemic subtypes only [16, 17]. CV mortality was defined as death due to coronary heart disease or cerebrovascular disease, and all-cause mortality was defined as death due to any cause. Outcome identification and ascertainment methods varied across the data sources (ESM). It was assumed that any CV effects due to the OAB medications would present shortly after new use and continue while patients received medication, and that the effects would decline after discontinuation.

2.6 Follow-Up Period

Follow-up for outcome occurrence started on the day after the start of the episode and ended at the earliest of: occurrence of the outcome, end of the study period, last date with validated adverse CV or death outcomes, disenrolment from the data source (e.g. emigration, death), prescription of non-tablet antimuscarinics or a prescription for multiple OAB medications on the same day. Patients were eligible to contribute additional treatment episodes until the earliest of the end of the study period, last date of data with validated adverse CV or death outcomes or disenrolment from the data source.

Once an outcome occurred, patients were no longer followed for subsequent occurrences of the same outcome. Patients could experience multiple outcome types. However, the occurrence of CV or all-cause mortality censored the treatment episode and the study follow-up. For MACE, the treatment episode and the follow-up were censored at the date of the first targeted CV event.

2.7 Statistical Methods

Data were analysed separately for the five data sources and all coding was done independently by each partner. A meta-analysis was then conducted to combine the results.

Within each population, PSs were estimated for each eligible treatment episode by modelling the probability for treatment with mirabegron or antimuscarinics, conditional on the baseline covariates. A unique calendar year-specific PS was calculated for each eligible episode using relevant baseline variables from the 12 months prior to the start of each episode. In the USA, mirabegron was approved in June 2012, thus episodes from 2012 and 2013 were combined. The PS model was built using pre-specified variables from all databases (e.g. age group, prior OAB medication use), database-specific variables (e.g. length of health plan enrolment) and database-specific empirically defined variables (e.g. most frequent International Classification of Diseases diagnosis codes). Several checks were conducted to support

the variable selection process, including correlations among pre-defined covariates and empirically defined variables and review of the PS distributions.

The mirabegron episodes were matched to the antimuscarinic episodes (1:1 ratio) based on the PSs using a greedy matching algorithm [18] with a calliper of 0.009. Matches were restricted to episodes from the same calendar year and patients of the same sex and age category (< 65 years, ≥ 65 years). If a patient contributed more than one antimuscarinic episode (even if the drugs were different), these episodes were joined together to form a single period of current exposure to antimuscarinics. This only occurred if the prescription date for the second treatment episode was within the grace period for the first episode.

Absolute standardised mean differences were calculated to assess the final PS model balance [19]. Variables with a difference ≤ 0.1 were considered balanced. If a variable had a difference of > 0.1 to ≤ 0.2, further assessment was required. A difference > 0.2 was defined as unbalanced and required further PS model modifications (e.g. interaction terms), or adjustment for the variable in the outcome model.

For the matched episodes, current exposure duration and the number of prescriptions per episode of current exposure were analysed, along with the number of matched episodes and episode type (mirabegron or antimuscarinics) contributed per patient. Baseline characteristics for the mirabegron and individual antimuscarinic episodes were evaluated before and after matching (data before matching not shown). Descriptive analyses were conducted using summary measures (frequencies, proportions, medians and interquartile ranges).

For each outcome, incidence rates per 1000 person-years and 95% confidence intervals (CIs) were calculated. Cox proportional hazard models were used to estimate HRs to compare outcome incidence rates between the mirabegron and the antimuscarinic episodes.

The incidence rate of each outcome within current exposure of the mirabegron episodes was compared with the antimuscarinic episodes as a group. This analysis was also conducted by age category, OAB medication use prior to the start of the treatment episode (naïve, non-naïve) and restricting episodes to those at high risk for CV events. Non-naïve users were defined as patients who received a prescription for another OAB medication during the prior 12 months. Episodes were defined as at high risk for CV events if the patient had at least one prior diagnosis for stroke, transient ischaemic attack, coronary artery disease, angina, AMI, heart failure, cardiac arrhythmias, chronic renal insufficiency or peripheral arterial or vascular disease or at least one prior percutaneous transluminal coronary angioplasty, coronary artery bypass graft or carotid endarterectomy. Episodes from patients with three or more of the following, hyperlipidaemia, hypertension, diabetes mellitus or age ≥ 80 years, were

also considered to be at high risk. Secondary analyses were conducted that separated results according to sex and a prior history of AMI or stroke; investigated the potential effect of residual confounding by alcohol/substance abuse, obesity and tobacco smoking for each data source and censored all adverse CV outcomes upon occurrence of the first individual event. For the residual confounding analysis, assessments were performed prior to pooling estimates across data sources using a spreadsheet based on previously published formula and approaches [20].

Pooled HR estimates across data sources were generated using Comprehensive Meta-Analysis software (RRID:SCR_012779) Version 3.0 [21]. Given anticipated heterogeneity in episode characteristics, prescribing patterns and the availability of covariate information across data sources, both random-effects and fixed-effects models were implemented, with point estimates representing a weighted HR of the results from the individual data sources using an inverse-variance-weighting approach [22]. Heterogeneity across data sources was assessed using the I^2 test, with $I^2 > 50%$ used to indicate substantial heterogeneity. Individual study population data were also generated, and the HRs were calculated using Cox proportional hazard models and the corresponding Wald-based CIs were derived from these models.

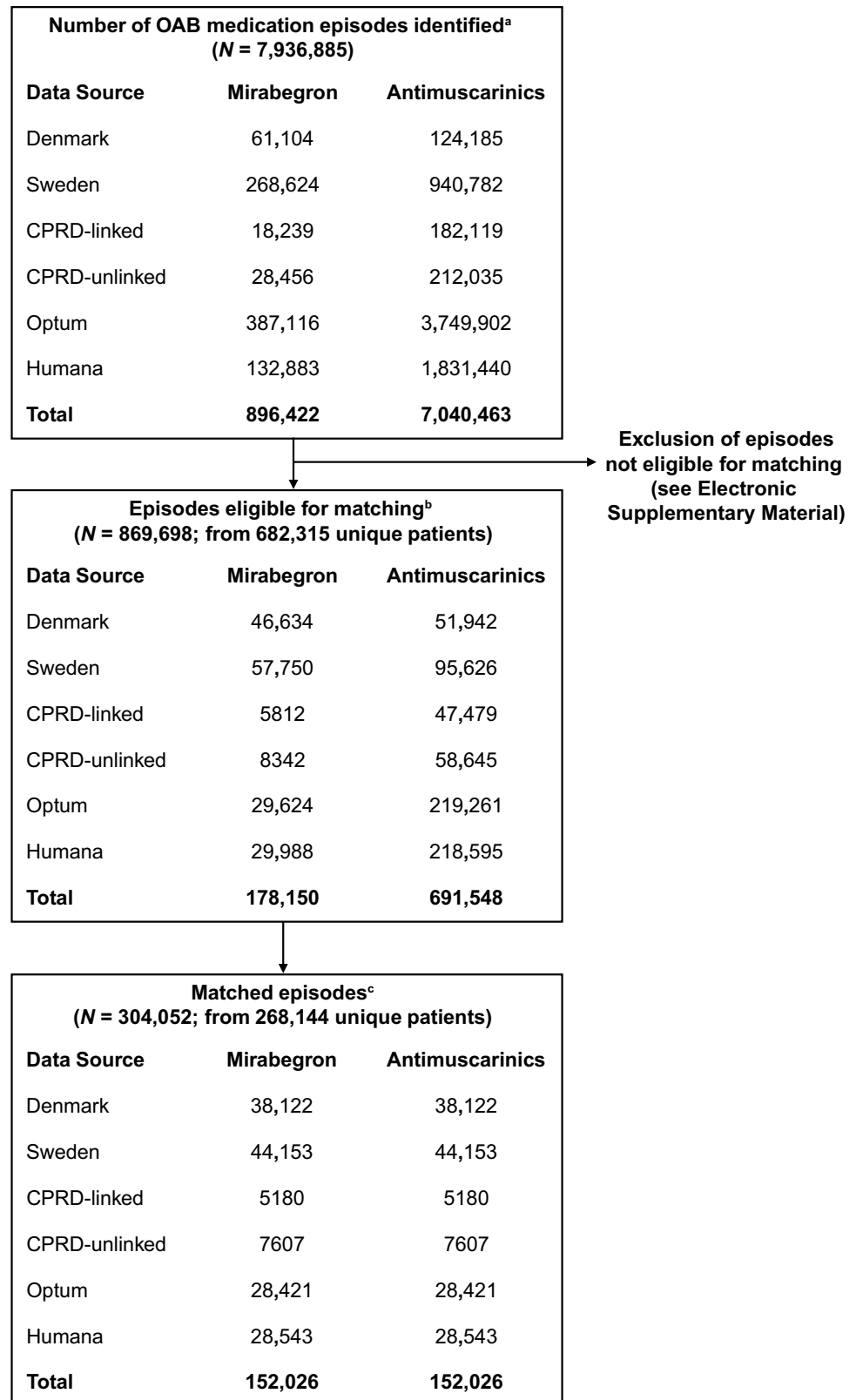
3 Results

3.1 Treatment Episodes

A total of 896,422 mirabegron episodes and 7,040,463 antimuscarinic episodes were identified prior to applying the inclusion and exclusion criteria (Fig. 1). Of these, 178,150 mirabegron and 691,548 antimuscarinic episodes were eligible for PS matching. A total of 152,026 mirabegron and 152,026 antimuscarinic episodes were subsequently matched one-to-one; the reasons for exclusion of the episodes are shown in the ESM. Furthermore, no episodes of oxybutynin use were identified in Denmark.

The median duration of current exposure for the matched treatment episodes ranged from 67 days (Humana database) to 165 days (Danish data) for mirabegron, and from 45 days (CPRD-linked) to 114 days (Danish data) for antimuscarinics. Median length of episode follow-up for the matched treatment episodes ranged from 173 days (ORD) to 564 days (Danish data) for mirabegron, and from 183 days (ORD) to 401 days (Danish data) for antimuscarinics. The median number of prescriptions per episode of current exposure ranged from one (Swedish data and Humana) to two (Danish data, CPRD-linked, CPRD-unlinked and ORD) for mirabegron, and also from one (Danish data, Swedish

Fig. 1 Matched cohort creation. *CPRD* Clinical Practice Research Datalink, *OAB* overactive bladder. ^aEach OAB medication prescription was an episode considered for inclusion in the study. ^bEpisodes eligible for matching were those with prescriptions of mirabegron or antimuscarinics that met all study inclusion criteria. ^cMatched episodes were those that met all study inclusion/exclusion criteria and were propensity score matched using a 1:1 ratio



data, CPRD-linked, CPRD-unlinked and Humana) to two (ORD) for antimuscarinics.

Within each data source, the majority of the 268,144 total patients contributed only one matched episode, ranging from 79.6% of patients from the Danish register to 92.7% of patients from the Swedish register. Of the 31,231 total patients who contributed two or more matched episodes, approximately 60% contributed both mirabegron and antimuscarinic episodes (from 44.8% of patients from the Swedish register to 69.5% of patients from the CPRD-linked database).

3.2 Baseline Characteristics

Strong balance was achieved for each data source across a range of measured baseline covariates (Table 1). Owing to sex and age group category matching, identical data were generated for these parameters in the mirabegron and antimuscarinic cohorts. Overall, 72.6% (220,798) and 63.1% (192,006) of the episodes were from patients who were ≥ 65 years old and female, respectively. Although differences in CV disease risk factors were observed across data sources, the matched mirabegron and antimuscarinic episodes within each data source were well balanced with respect to these risk factors.

3.3 CV and Mortality Outcomes

The results below are from pooled analyses unless otherwise stated. As the findings were usually similar with the fixed-effects and the random-effects models, the results shown are primarily from the fixed-effects model. When substantial heterogeneity was observed across data sources, results from both models are included.

The incidence rates per 1000 person-years during current mirabegron exposure were 21.99 (95% CI 20.97–23.05) for MACE, 5.12 (95% CI 4.63–5.64) for AMI, 12.78 (95% CI 12.00–13.59) for stroke, 5.67 (95% CI 5.17–6.22) for CV mortality and 34.95 (95% CI 33.66–36.28) for all-cause mortality (Table 2). During current antimuscarinic exposure, the incidence rates were 22.71 (95% CI 21.59–23.86) for MACE, 5.32 (95% CI 4.80–5.89) for AMI, 12.05 (95% CI 11.25–12.90) for stroke, 7.09 (95% CI 6.48–7.74) for CV mortality and 41.10 (95% CI 39.60–42.65) for all-cause mortality.

The incidence rates of MACE, AMI and stroke were similar among current users of mirabegron relative to current users of antimuscarinics, with HRs close to 1 and the 95% CIs encompassing 1 (MACE: HR 0.94, 95% CI 0.88–1.01; AMI: HR 0.94, 95% CI 0.81–1.08; stroke: HR 1.00, 95% CI 0.91–1.10). The incidence rates of CV mortality and all-cause mortality were no higher with mirabegron compared with antimuscarinics, with HRs of 0.83 (95% CI 0.73–0.95)

and 0.80 (95% CI 0.76–0.84), respectively. The HRs for each outcome in the individual study populations (non-pooled data) were generally similar to those in the combined population (Fig. 2 and ESM).

In the analysis according to age category, the HRs for MACE, AMI and stroke ranged from 0.74 to 1.04 (< 65 years) and from 0.94 to 0.98 (≥ 65 years), and the corresponding 95% CIs all encompassed 1 (Table 3). Among those aged < 65 years, the HR for all-cause mortality was 0.76 (95% CI 0.62–0.93). In contrast to the overall results, the 95% CI was relatively large and encompassed 1 for the HR for CV mortality in this age category (HR 0.92, 95% CI 0.53–1.59). Among those aged ≥ 65 years, the HRs for CV mortality and all-cause mortality displayed the same associations as the overall results.

For the analysis by prior OAB medication use, results showed that the HRs among non-naïve OAB medication users ranged from 0.97 to 1.09 and the 95% CIs encompassed 1 (Table 4). In naïve users, the HRs for MACE, AMI and stroke ranged from 0.89 to 0.96; the 95% CIs for AMI and stroke encompassed 1, whereas the upper limit of the CI for the MACE estimate was slightly less than 1 (HR 0.90, 95% CI 0.83–0.97). The HR for all-cause mortality for naïve users was 0.73 (95% CI 0.69–0.78). Substantial heterogeneity was observed across the data sources in terms of CV mortality for naïve users ($I^2 = 62\%$). The HR for CV mortality was 0.79 (95% CI 0.68–0.92) using the fixed-effects model and 0.75 (95% CI 0.56–1.02) using the random-effects model.

Approximately 70% of the mirabegron and antimuscarinic episodes were considered at high risk for CV events (Table 5). When the data were restricted to these episodes, similar findings to the overall results were observed.

For the analysis by sex, the HRs among female patients were similar to the overall results (ESM). The results among male patients were also generally similar to the overall results. However, substantial heterogeneity among male patients was observed across the data sources for both MACE ($I^2 = 59\%$) and CV mortality ($I^2 = 53\%$). For male patients, the HRs for the fixed-effects model were 0.95 (95% CI 0.86–1.04) for MACE and 0.91 (95% CI 0.76–1.08) for CV mortality. With the random-effects model, the HRs were 0.92 (95% CI 0.76–1.12) for MACE and 0.89 (95% CI 0.65–1.21) for CV mortality.

For the analysis by prior history of AMI or stroke, similar results were observed to the overall findings, with the exception of CV mortality (ESM). For patients with a prior history of AMI or stroke, the 95% CIs for CV mortality encompassed 1 (HR 0.84, 95% CI 0.69–1.03) and for patients without a prior history, the upper confidence limit was equal to 1 (HR 0.84, 95% CI 0.71–1.00). In addition, substantial heterogeneity was observed for CV mortality in patients who had a prior history of AMI or stroke ($I^2 = 59\%$) and the HR

Table 1 Baseline demographics and cardiovascular disease risk factors for the matched mirabegron and antimuscarinic episodes

Characteristic	Denmark (n = 38,122)	Sweden (n = 44,153)	CPRD-linked (n = 5180)	CPRD-unlinked (n = 7607)	Optum (n = 28,421)	Humana (n = 28,543)	Total (n = 152,026)
Age categories							
18 to ≤ 44 years	2856 (7.5)	3203 (7.3)	500 (9.7)	712 (9.4)	541 (1.9)	688 (2.4)	8500 (5.6)
45 to ≤ 54 years	3503 (9.2)	3898 (8.8)	679 (13.1)	1000 (13.1)	1353 (4.8)	1430 (5.0)	11,863 (7.8)
55 to ≤ 64 years	5590 (14.7)	6751 (15.3)	831 (16.0)	1378 (18.1)	3265 (11.5)	3449 (12.1)	21,264 (14.0)
65 to ≤ 74 years	11,700 (30.7)	14,824 (33.6)	1358 (26.2)	1971 (25.9)	10,165 (35.8)	10,132 (35.5)	50,150 (33.0)
≥ 75 years	14,473 (38.0)	15,477 (35.1)	1812 (35.0)	2546 (33.5)	13,097 (46.1)	12,844 (45.0)	60,249 (39.6)
Age ≥ 65 years ^a	26,173 (68.7)	30,301 (68.6)	3170 (61.2)	4517 (59.4)	23,262 (81.8)	22,976 (80.5)	110,399 (72.6)
Female ^a	21,018 (55.1)	26,408 (59.8)	3532 (68.2)	5195 (68.3)	20,156 (70.9)	19,694 (69.0)	96,003 (63.1)
Oestrogen use ^b	10,476 (49.8)	14,890 (56.4)	479 (13.6)	635 (12.2)	283 (1.4)	1340 (6.8)	28,103 (29.3)
Hypertension	25,986 (68.2)	27,031 (61.2)	3180 (61.4)	4895 (64.3)	26,229 (92.3)	25,211 (88.3)	112,532 (74.0)
Stroke	1961 (5.1)	3452 (7.8)	229 (4.4)	340 (4.5)	2695 (9.5)	4040 (14.2)	12,717 (8.4)
AMI	746 (2.0)	3085 (7.0)	228 (4.4)	301 (4.0)	1047 (3.7)	1662 (5.8)	7069 (4.6)
Diabetes mellitus	5801 (15.2)	5778 (13.1)	876 (16.9)	1183 (15.6)	11,747 (41.3)	13,891 (48.7)	39,276 (25.8)
Hyperlipidaemia	16,959 (44.5)	17,043 (38.6)	2406 (46.4)	3659 (48.1)	20,792 (73.2)	25,148 (88.1)	86,007 (56.6)
Antiplatelet medication use	15,660 (41.1)	20,730 (47.0)	478 (9.2)	909 (11.9)	3708 (13.0)	4373 (15.3)	45,858 (30.2)
Previous use of OAB medications^c							
Mirabegron	0 (0.0)	0 (0.0)	0 (0.0)	0 (0.0)	0 (0.0)	0 (0.0)	0 (0.0)
Any antimuscarinic medication	6134 (16.1)	6861 (15.5)	3164 (61.1)	4847 (63.7)	8515 (30.0)	11,486 (40.2)	41,007 (27.0)
Solifenacin	2450 (6.4)	2620 (5.9)	1770 (34.2)	2867 (37.7)	3242 (11.4)	3352 (11.7)	16,301 (10.7)
Oxybutynin	24 (0.1)	288 (0.7)	860 (16.6)	1170 (15.4)	4369 (15.4)	5959 (20.9)	12,670 (8.3)
Tolterodine	1620 (4.2)	2595 (5.9)	738 (14.2)	1024 (13.5)	386 (1.4)	2024 (7.1)	8387 (5.5)
Fesoterodine	1115 (2.9)	1284 (2.9)	394 (7.6)	603 (7.9)	1064 (3.7)	962 (3.4)	5422 (3.6)
Trospium	1190 (3.1)	N/A ^d	395 (7.6)	456 (6.0)	393 (1.4)	512 (1.8)	2946 (1.9)
Darifenacin	221 (0.6)	408 (0.9)	49 (0.9)	112 (1.5)	262 (0.9)	349 (1.2)	1401 (0.9)
Number of previous treatment episodes, mean (SD)^e							
Mirabegron	0.10 (0.40)	0.07 (0.47)	0.04 (0.25)	0.07 (0.39)	1.10 (3.16)	0.11 (0.92)	NC
Any antimuscarinic medication	0.30 (0.80)	2.48 (7.15)	2.89 (4.07)	2.76 (3.49)	3.39 (7.74)	4.96 (8.64)	NC
Solifenacin	0.10 (0.40)	0.97 (4.13)	1.22 (2.26)	1.26 (1.99)	0.99 (4.02)	1.40 (4.59)	NC
Oxybutynin	0.00 (0.00)	0.05 (0.96)	0.73 (2.49)	0.60 (1.93)	1.28 (4.65)	2.13 (5.47)	NC
Tolterodine	0.10 (0.40)	0.86 (4.92)	0.47 (1.50)	0.43 (1.41)	0.19 (1.71)	0.76 (3.55)	NC
Fesoterodine	0.00 (0.40)	0.45 (2.35)	0.24 (0.92)	0.25 (0.89)	0.35 (2.38)	0.31 (2.11)	NC
Trospium	0.00 (0.30)	N/A ^d	0.22 (1.04)	0.20 (0.90)	0.13 (1.53)	0.17 (1.69)	NC
Darifenacin	0.00 (0.10)	0.16 (1.40)	0.02 (0.21)	0.03 (0.27)	0.15 (1.53)	0.19 (1.74)	NC
Healthcare utilisation^f							
Hospitalisations ^g	2486 (6.5)	1720 (3.9)	357 (6.9)	500 (6.6)	1239 (4.4)	746 (2.6)	7048 (4.6)
Number of hospitalisations, mean (SD)	0.80 (1.80)	0.46 (1.10)	0.96 (2.25)	0.66 (1.38)	0.37 (0.81)	0.33 (0.76)	NC
Number of ER visits, mean (SD)	0.10 (0.40)	N/A ^h	0.41 (0.97)	0.31 (0.85)	1.09 (2.30)	1.28 (2.78)	NC
Number of outpatient physician office visits, mean (SD)	2.90 (3.00)	3.79 (4.35)	11.48 (8.60)	10.81 (7.82)	15.27 (11.13)	13.25 (8.48)	NC
Number of medications, mean (SD)	9.10 (5.90)	10.93 (6.49)	11.13 (6.25)	11.98 (6.27)	15.66 (7.98)	15.69 (7.58)	NC
Antimuscarinics							
	(n = 38,122)	(n = 44,153)	(n = 5180)	(n = 7607)	(n = 28,421)	(n = 28,543)	(n = 152,026)
Age categories							
18 to ≤ 44 years	2808 (7.4)	3225 (7.3)	467 (9.0)	754 (9.9)	544 (1.9)	611 (2.1)	8409 (5.5)
45 to ≤ 54 years	3446 (9.0)	3628 (8.2)	656 (12.7)	1004 (13.2)	1369 (4.8)	1395 (4.9)	11,498 (7.6)
55 to ≤ 64 years	5695 (14.9)	6999 (15.9)	887 (17.1)	1332 (17.5)	3246 (11.4)	3561 (12.5)	21,720 (14.3)
65 to ≤ 74 years	11,748 (30.8)	14,085 (31.9)	1320 (25.5)	1952 (25.7)	10,260 (36.1)	10,135 (35.5)	49,500 (32.6)

Table 1 (continued)

Antimuscarinics	(n = 38,122)	(n = 44,153)	(n = 5180)	(n = 7607)	(n = 28,421)	(n = 28,543)	(n = 152,026)
≥ 75 years	14,425 (37.8)	16,216 (36.7)	1850 (35.7)	2565 (33.7)	13,002 (45.8)	12,841 (45.0)	60,899 (40.1)
Age ≥ 65 years ^a	26,173 (68.7)	30,301 (68.6)	3170 (61.2)	4517 (59.4)	23,262 (81.8)	22,976 (80.5)	110,399 (72.6)
Female ^a	21,018 (55.1)	26,408 (59.8)	3532 (68.2)	5195 (68.3)	20,156 (70.9)	19,694 (69.0)	96,003 (63.1)
Oestrogen use ^b	10,329 (49.1)	15,296 (57.9)	442 (12.5)	576 (11.1)	296 (1.5)	1372 (7.0)	28,311 (29.5)
Hypertension	25,982 (68.2)	27,109 (61.4)	3208 (61.9)	4973 (65.4)	26,189 (92.1)	25,187 (88.2)	112,648 (74.1)
Stroke	1984 (5.2)	3497 (7.9)	247 (4.8)	347 (4.6)	2716 (9.6)	4076 (14.3)	12,867 (8.5)
AMI	769 (2.0)	3160 (7.2)	246 (4.7)	346 (4.5)	1060 (3.7)	1676 (5.9)	7257 (4.8)
Diabetes mellitus	5806 (15.2)	5813 (13.2)	882 (17.0)	1304 (17.1)	11,741 (41.3)	13,980 (49.0)	39,526 (26.0)
Hyperlipidaemia	17,029 (44.7)	17,082 (38.7)	2414 (46.6)	3691 (48.5)	20,733 (72.9)	25,241 (88.4)	86,190 (56.7)
Antiplatelet medication use	15,660 (41.1)	20,781 (47.1)	457 (8.8)	880 (11.6)	3680 (12.9)	4462 (15.6)	45,920 (30.2)
Previous use of OAB medications ^c							
Mirabegron	5531 (14.5)	0 (0.0)	402 (7.8)	561 (7.4)	1971 (6.9)	10 (< 0.1)	8475 (5.6)
Any antimuscarinic medication	5891 (15.5)	6572 (14.9)	3174 (61.3)	4845 (63.7)	8066 (28.4)	11,358 (39.8)	39,906 (26.2)
Solifenacin	2268 (5.9)	2362 (5.3)	1807 (34.9)	2846 (37.4)	3103 (10.9)	3342 (11.7)	15,728 (10.3)
Oxybutynin	15 (< 0.1)	275 (0.6)	803 (15.5)	1093 (14.4)	3936 (13.9)	5667 (19.9)	11,789 (7.8)
Tolterodine	1588 (4.2)	2569 (5.8)	694 (13.4)	992 (13.0)	402 (1.4)	2053 (7.2)	8298 (5.5)
Fesoterodine	1056 (2.8)	1206 (2.7)	307 (5.9)	514 (6.8)	993 (3.5)	920 (3.2)	4996 (3.3)
Trospium	1177 (3.1)	N/A ^d	290 (5.6)	383 (5.0)	382 (1.3)	460 (1.6)	2692 (1.8)
Darifenacin	207 (0.5)	407 (0.9)	25 (0.5)	64 (0.8)	257 (0.9)	327 (1.1)	1287 (0.8)
Number of previous treatment episodes, mean (SD) ^e							
Mirabegron	0.20 (0.50)	0.06 (0.45)	0.14 (0.55)	0.15 (0.65)	0.26 (1.61)	0.09 (0.74)	NC
Any antimuscarinic medication	0.40 (0.90)	2.51 (6.30)	2.69 (3.52)	2.74 (3.56)	4.05 (7.62)	4.74 (8.31)	NC
Solifenacin	0.20 (0.60)	0.96 (3.77)	1.19 (2.10)	1.30 (2.10)	1.34 (4.45)	1.41 (4.70)	NC
Oxybutynin	0.00 (0.00)	0.07 (1.03)	0.66 (2.05)	0.58 (1.94)	1.81 (4.77)	1.99 (5.19)	NC
Tolterodine	0.10 (0.50)	0.87 (3.87)	0.45 (1.41)	0.42 (1.40)	0.21 (1.78)	0.77 (3.38)	NC
Fesoterodine	0.10 (0.40)	0.44 (2.46)	0.20 (0.74)	0.24 (1.04)	0.33 (2.40)	0.28 (2.06)	NC
Trospium	0.10 (0.30)	N/A ^d	0.18 (0.75)	0.17 (0.78)	0.09 (1.27)	0.13 (1.43)	NC
Darifenacin	0.00 (0.20)	0.17 (1.47)	0.02 (0.19)	0.02 (0.23)	0.15 (1.62)	0.16 (1.56)	NC
Healthcare utilisation ^f							
Hospitalisations ^g	2910 (7.6)	1677 (3.8)	361 (7.0)	397 (5.2)	1199 (4.2)	648 (2.3)	7192 (4.7)
Number of hospitalisations, mean (SD)	0.90 (1.90)	0.46 (1.08)	0.95 (2.47)	0.66 (2.47)	0.37 (0.82)	0.33 (0.75)	NC
Number of ER visits, mean (SD)	0.10 (0.40)	N/A ^h	0.43 (1.04)	0.32 (0.90)	1.12 (2.33)	1.31 (2.54)	NC
Number of outpatient physician office visits, mean (SD)	2.80 (3.00)	3.84 (5.48)	11.53 (9.09)	11.00 (8.29)	15.44 (11.43)	12.96 (8.62)	NC
Number of medications, mean (SD)	9.20 (6.00)	10.97 (6.65)	11.16 (6.24)	12.11 (6.46)	15.72 (7.89)	15.75 (7.59)	NC

AMI acute myocardial infarction, CPRD Clinical Practice Research Datalink, ER emergency room, N/A not applicable, NC not calculated, OAB overactive bladder, SD standard deviation

Data shown in terms of n (%), unless otherwise specified

^aThe age and sex distributions for the mirabegron and antimuscarinic groups are identical, as the episodes were matched in a 1:1 ratio using sex and age group (< 65 years, ≥ 65 years) categories

^bPercentages were calculated using the number of episodes from female patients as the denominator

^cIncludes all prescriptions that occurred in the 12 months prior to each episode initiation date

^dTrospium was not available for use in Sweden during the study period

^eIncludes all previous treatment episodes that occurred up to 5 years prior to each episode initiation date

^fIn the past 12 months

^gAny admissions in the past 1 month

^hHealthcare contacts are classified as planned or unplanned in the Swedish National Registers, therefore ER visits were not available

Table 2 Adverse CV and mortality outcomes during current use of mirabegron or antimuscarinics (matched episodes)

Outcomes	Episodes ^a	Cases	Person-years	Incidence rate (95% CI) ^b	Fixed effects HR (95% CI) ^c	Random effects HR (95% CI) ^c	<i>I</i> ² (%)
MACE^d							
Mirabegron	151,481	1707	77,618	21.99 (20.97–23.05)	0.94 (0.88–1.01)	0.94 (0.84–1.04)	37.96
Antimuscarinics	147,859	1536	67,645	22.71 (21.59–23.86)	Ref	Ref	
AMI							
Mirabegron	151,708	401	78,384	5.12 (4.63–5.64)	0.94 (0.81–1.08)	0.94 (0.81–1.08)	0.00
Antimuscarinics	148,127	363	68,203	5.32 (4.80–5.89)	Ref	Ref	
Stroke							
Mirabegron	151,560	995	77,886	12.78 (12.00–13.59)	1.00 (0.91–1.10)	1.02 (0.88–1.17)	33.81
Antimuscarinics	147,965	818	67,877	12.05 (11.25–12.90)	Ref	Ref	
CV mortality							
Mirabegron	151,798	446	78,602	5.67 (5.17–6.22)	0.83 (0.73–0.95)	0.83 (0.71–0.97)	23.47
Antimuscarinics	148,244	485	68,402	7.09 (6.48–7.74)	Ref	Ref	
All-cause mortality							
Mirabegron	151,798	2736	78,279	34.95 (33.66–36.28)	0.80 (0.76–0.84)	0.80 (0.76–0.84)	0.08
Antimuscarinics	148,244	2802	68,172	41.10 (39.60–42.65)	Ref	Ref	

AMI acute myocardial infarction, CI confidence interval, CMA Comprehensive Meta-Analysis, CV cardiovascular, HR hazard ratio, MACE major adverse cardiovascular events, Ref reference group

^aWhere multiple individual antimuscarinic treatment episodes were joined to form a single period of current exposure, only the first matched episode during the period of current exposure was counted

^bIncidence rates were calculated in terms of per 1000 person-years. CIs were calculated using Byar's formula [23]

^cHRs and corresponding CIs were calculated using CMA Version 3 software

^dThe number of AMI, stroke and CV mortality cases do not sum to the number of MACE cases. For patients with more than one of the individual outcomes, the earliest individual outcome was counted towards the MACE total

for the random-effects model was 0.80 (95% CI 0.57–1.14). Substantial heterogeneity was also observed for MACE in patients without a prior history of AMI or stroke ($I^2 = 53\%$). The HR for MACE was 0.97 (95% CI 0.88–1.06) using the fixed-effects model and 0.96 (95% CI 0.82–1.12) using the random-effects model.

The residual confounding analysis indicated that the presence of bias was < 9% for smoking, alcohol/substance abuse and obesity in each of the data sources (ESM). When the data were censored for all CV outcomes upon occurrence of the first event, the results were nearly identical to those generated in the main analysis (ESM).

4 Discussion

This large, multinational study showed no appreciable differences in the incidences of MACE, AMI and stroke between current users of mirabegron and current users of antimuscarinics. Furthermore, the risks of CV mortality and all-cause mortality were no higher with mirabegron compared

with antimuscarinics. These findings were consistently noted across the subgroup analyses.

In advance of this study, CV validation studies were conducted for each data source to confirm outcome definitions and to assess the suitability of each data source for the post-authorisation safety programme (results published for the studies conducted using the CPRD and Danish register data [24–26], Swedish register data were presented at the International Conference on Pharmacoepidemiology and Therapeutic Risk Management in 2016 [27], results unpublished for the ORD and Humana database studies). As mirabegron exposure results were not widely available at the time these validation studies were initiated, the investigations were conducted using data from antimuscarinic users only. Although specific associations were observed in the analyses conducted for some of the data sources [24, 27], no clear and consistent trends in the incidence rate ratios for adverse CV outcomes were observed following exposure to a range of antimuscarinics. In view of these observations, as well as the evidence that patient characteristics were similar in individuals who initiated each of the antimuscarinics, it was decided to undertake the primary analysis for the

Fig. 2 Adverse CV and mortality outcomes during current use of mirabegron or antimuscarinics (matched episodes), by individual data source. *AMI* acute myocardial infarction, *CI* confidence interval, *CPRD* Clinical Practice Research Datalink, *CV* cardiovascular, *HR* hazard ratio, *MACE* major adverse cardiovascular events. HRs were calculated using Cox proportional hazard models and the corresponding Wald-based CIs were generated from the Cox models

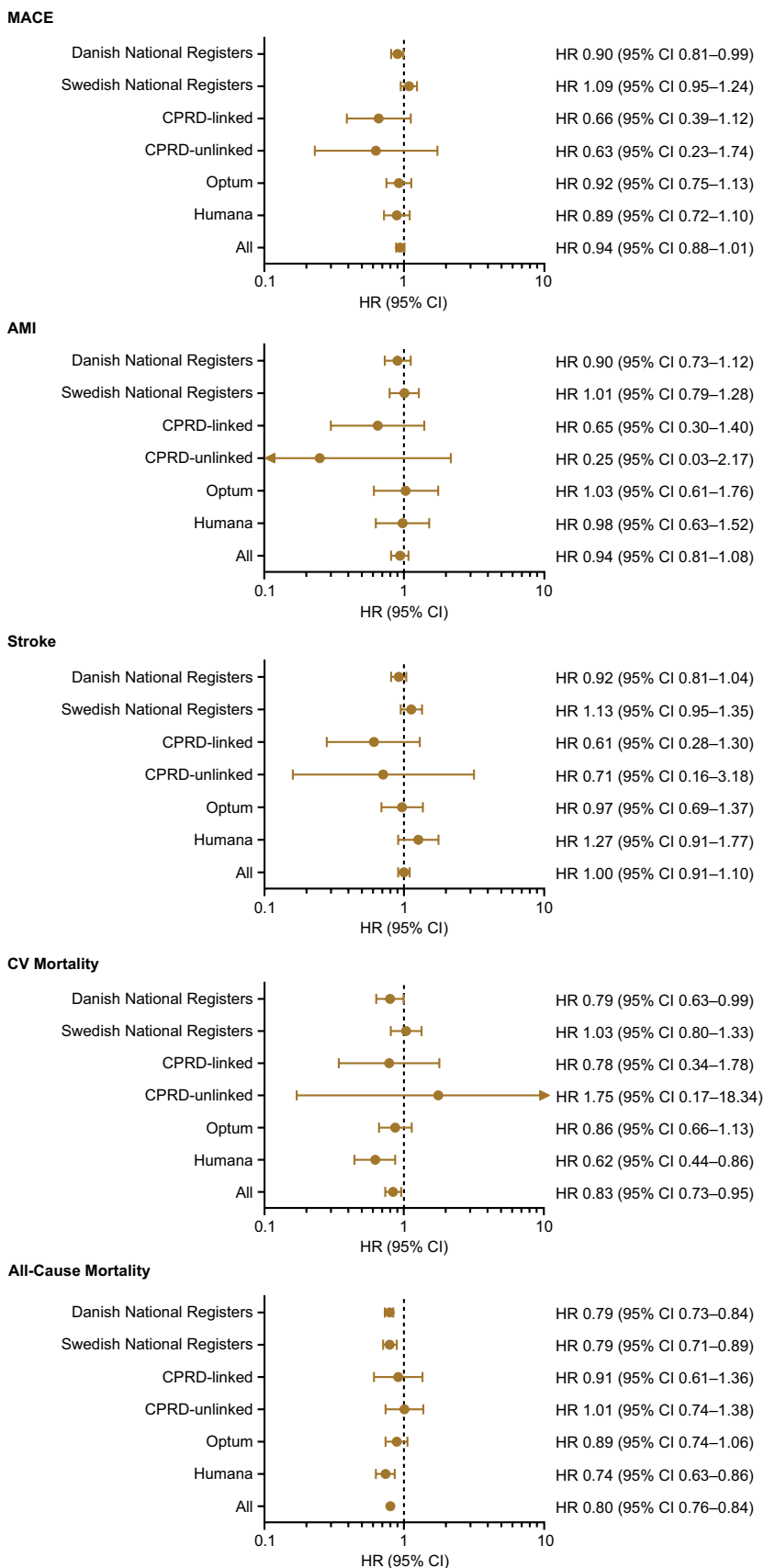


Table 3 Adverse CV and mortality outcomes during current use of mirabegron or antimuscarinics (matched episodes), according to age group

Outcomes	Episodes ^a	Cases	Person-years	Incidence rate (95% CI) ^b	Fixed effects HR (95% CI) ^c	Random effects HR (95% CI) ^c	<i>I</i> ² (%)
Age < 65 years							
MACE ^d							
Mirabegron	41,534	180	20,930	8.60 (7.41–9.93)	0.95 (0.77–1.18)	0.97 (0.76–1.25)	10.86
Antimuscarinics	40,570	162	18,100	8.95 (7.65–10.41)	Ref	Ref	
AMI							
Mirabegron	41,554	41	21,036	1.95 (1.42–2.62)	0.74 (0.49–1.11)	0.74 (0.48–1.15)	6.49
Antimuscarinics	40,605	51	18,180	2.81 (2.11–3.66)	Ref	Ref	
Stroke							
Mirabegron	41,541	121	20,955	5.77 (4.81–6.87)	1.04 (0.78–1.37)	1.27 (0.73–2.20)	43.45
Antimuscarinics	40,581	91	18,126	5.02 (4.07–6.13)	Ref	Ref	
CV mortality							
Mirabegron	41,561	26	21,059	1.24 (0.83–1.78)	0.92 (0.53–1.59)	0.92 (0.53–1.59)	0.00
Antimuscarinics	40,616	26	18,204	1.43 (0.96–2.06)	Ref	Ref	
All-cause mortality							
Mirabegron	41,561	195	21,033	9.27 (8.04–10.64)	0.76 (0.62–0.93)	0.72 (0.54–0.96)	29.85
Antimuscarinics	40,616	215	18,186	11.82 (10.32–13.48)	Ref	Ref	
Age ≥ 65 years							
MACE ^d							
Mirabegron	109,947	1527	56,688	26.94 (25.61–28.31)	0.94 (0.87–1.01)	0.93 (0.84–1.03)	33.80
Antimuscarinics	107,289	1374	49,545	27.73 (26.30–29.23)	Ref	Ref	
AMI							
Mirabegron	110,154	360	57,347	6.28 (5.65–6.95)	0.97 (0.83–1.13)	0.97 (0.83–1.13)	0.00
Antimuscarinics	107,522	312	50,022	6.24 (5.57–6.96)	Ref	Ref	
Stroke							
Mirabegron	110,019	874	56,930	15.35 (14.36–16.40)	0.98 (0.89–1.09)	0.99 (0.86–1.13)	25.14
Antimuscarinics	107,384	727	49,752	14.61 (13.58–15.70)	Ref	Ref	
CV mortality							
Mirabegron	110,237	420	57,543	7.30 (6.63–8.02)	0.83 (0.73–0.95)	0.83 (0.71–0.96)	15.26
Antimuscarinics	107,628	459	50,198	9.14 (8.34–10.01)	Ref	Ref	
All-cause mortality							
Mirabegron	110,237	2541	57,246	44.39 (42.69–46.14)	0.80 (0.76–0.84)	0.81 (0.75–0.86)	16.04
Antimuscarinics	107,628	2587	49,985	51.76 (49.79–53.78)	Ref	Ref	

AMI acute myocardial infarction, CI confidence interval, CMA Comprehensive Meta-Analysis, CV cardiovascular, HR hazard ratio, MACE major adverse cardiovascular events, Ref reference group

^aWhere multiple individual antimuscarinic treatment episodes were joined to form a single period of current exposure, only the first matched episode during the period of current exposure was counted

^bIncidence rates were calculated in terms of per 1000 person-years. CIs were calculated using Byar's formula [23]

^cHRs and corresponding CIs were calculated using CMA Version 3 software

^dThe number of AMI, stroke and CV mortality cases do not sum to the number of MACE cases. For patients with more than one of the individual outcomes, the earliest individual outcome was counted towards the MACE total

current study using a comparison cohort that comprised all new users of antimuscarinics.

The MACE findings in our study are consistent with the findings of a systematic literature review conducted to examine the CV safety of β 3-adrenoreceptor agonists [28]. This review showed no increased risk of Antiplatelet Trialists' Collaboration/MACE events among patients treated with

mirabegron compared with patients treated with placebo (relative risk 0.24, 95% CI 0.02–1.69) in five 12-week trials [7–9, 28–30]. Similarly, a 1-year study included in the review found no evidence that the risk of these events was different for mirabegron vs tolterodine (relative risk 0.75, 95% CI 0.18–3.60) [28, 31]. However, this investigation lacked statistical power.

Table 4 Adverse CV and mortality outcomes during current use of mirabegron or antimuscarinics (matched episodes), according to prior OAB medication use

Outcomes	Episodes ^a	Cases	Person-years	Incidence rate (95% CI) ^b	Fixed effects HR (95% CI) ^c	Random effects HR (95% CI) ^c	I ² (%)
Naïve ^d							
MACE ^e							
Mirabegron	110,983	1336	57,581	23.20 (21.98–24.47)	0.90 (0.83–0.97)	0.89 (0.78–1.02)	44.54
Antimuscarinics	106,310	1143	45,302	25.23 (23.80–26.73)	Ref	Ref	
AMI							
Mirabegron	111,103	315	58,135	5.42 (4.85–6.04)	0.89 (0.76–1.05)	0.89 (0.70–1.13)	37.92
Antimuscarinics	106,444	273	45,678	5.98 (5.30–6.72)	Ref	Ref	
Stroke							
Mirabegron	111,031	782	57,780	13.53 (12.61–14.51)	0.96 (0.86–1.07)	0.98 (0.86–1.11)	15.60
Antimuscarinics	106,367	602	45,468	13.24 (12.21–14.33)	Ref	Ref	
CV mortality							
Mirabegron	111,158	338	58,293	5.80 (5.21–6.44)	0.79 (0.68–0.92)	0.75 (0.56–1.02)	62.40
Antimuscarinics	106,507	355	45,814	7.75 (6.97–8.59)	Ref	Ref	
All-cause mortality							
Mirabegron	111,158	2156	58,023	37.16 (35.61–38.75)	0.73 (0.69–0.78)	0.73 (0.68–0.80)	25.80
Antimuscarinics	106,507	2148	45,623	47.08 (45.12–49.10)	Ref	Ref	
Non-naïve ^d							
MACE ^e							
Mirabegron	40,498	371	20,037	18.52 (16.70–20.47)	1.07 (0.93–1.23)	1.07 (0.92–1.24)	2.92
Antimuscarinics	41,549	393	22,342	17.59 (15.92–19.39)	Ref	Ref	
AMI							
Mirabegron	40,605	86	20,249	4.25 (3.42–5.22)	1.09 (0.81–1.48)	1.09 (0.77–1.54)	16.93
Antimuscarinics	41,683	90	22,524	4.00 (3.23–4.89)	Ref	Ref	
Stroke							
Mirabegron	40,529	213	20,106	10.59 (9.24–12.09)	1.08 (0.89–1.31)	1.08 (0.86–1.36)	17.53
Antimuscarinics	41,598	216	22,409	9.64 (8.42–10.99)	Ref	Ref	
CV mortality							
Mirabegron	40,640	108	20,309	5.32 (4.38–6.39)	0.97 (0.75–1.26)	0.97 (0.75–1.26)	0.00
Antimuscarinics	41,737	130	22,588	5.76 (4.83–6.81)	Ref	Ref	
All-cause mortality							
Mirabegron	40,640	580	20,256	28.63 (26.37–31.04)	0.97 (0.87–1.09)	0.97 (0.87–1.09)	0.00
Antimuscarinics	41,737	654	22,549	29.00 (26.85–31.29)	Ref	Ref	

AMI acute myocardial infarction, CI confidence interval, CMA Comprehensive Meta-Analysis, CV cardiovascular, HR hazard ratio, MACE major adverse cardiovascular events, OAB overactive bladder, Ref reference group

^aWhere multiple individual antimuscarinic treatment episodes were joined to form a single period of current exposure, only the first matched episode during the period of current exposure was counted

^bIncidence rates were calculated in terms of per 1000 person-years. CIs were calculated using Byar's formula [23]

^cHRs and corresponding CIs were calculated using CMA Version 3 software

^dNon-naïve users had received a prescription for another OAB medication during the prior 12 months and naïve users had not received a prescription for another OAB medication during the prior 12 months

^eThe number of AMI, stroke and CV mortality cases do not sum to the number of MACE cases. For patients with more than one of the individual outcomes, the earliest individual outcome was counted towards the MACE total

The findings for AMI and stroke are consistent with previous similar non-interventional studies using real-world data. An analysis conducted by the Food and Drug Administration from November 2012 to May 2014 (the Mini-Sentinel

programme) involved investigating diagnoses of acute AMI and stroke events in matched new users of mirabegron or oxybutynin [32]. In agreement with our findings, the Mini-Sentinel programme found no evidence for differences in the

Table 5 Adverse CV and mortality outcomes during current use of mirabegron or antimuscarinics (matched episodes), restricted to episodes at high risk for CV events

Outcomes	Episodes ^{a,b}	Cases	Person-years	Incidence rate (95% CI) ^c	Fixed effects HR (95% CI) ^d	Random effects HR (95% CI) ^d	<i>I</i> ² (%)
MACE^e							
Mirabegron	103,046	1526	54,300	28.10 (26.72–29.54)	0.95 (0.88–1.02)	0.95 (0.85–1.05)	32.56
Antimuscarinics	100,726	1365	47,172	28.94 (27.43–30.50)	Ref	Ref	
AMI							
Mirabegron	103,263	349	54,994	6.35 (5.71–7.04)	0.93 (0.80–1.09)	0.93 (0.80–1.09)	0.00
Antimuscarinics	100,984	317	47,681	6.65 (5.95–7.41)	Ref	Ref	
Stroke							
Mirabegron	103,119	885	54,542	16.23 (15.18–17.32)	1.00 (0.91–1.11)	1.03 (0.89–1.20)	32.06
Antimuscarinics	100,830	722	47,378	15.24 (14.16–16.38)	Ref	Ref	
CV mortality							
Mirabegron	103,347	414	55,190	7.50 (6.81–8.25)	0.84 (0.74–0.96)	0.84 (0.72–0.97)	10.19
Antimuscarinics	101,099	447	47,856	9.34 (8.51–10.24)	Ref	Ref	
All-cause mortality							
Mirabegron	103,347	2458	54,897	44.78 (43.03–46.57)	0.82 (0.77–0.87)	0.82 (0.77–0.87)	0.00
Antimuscarinics	101,099	2460	47,654	51.62 (49.61–53.69)	Ref	Ref	

AMI acute myocardial infarction, CI confidence interval, CMA Comprehensive Meta-Analysis, CV cardiovascular, HR hazard ratio, MACE major adverse cardiovascular events, Ref reference group

^aWhere multiple individual antimuscarinic treatment episodes were joined to form a single period of current exposure, only the first matched episode during the period of current exposure was counted

^bEpisodes were defined as high risk if the patient had at least one prior diagnosis for stroke, transient ischaemic attack, coronary artery disease, angina, AMI, heart failure, cardiac arrhythmias, chronic renal insufficiency or peripheral arterial or vascular disease or at least one prior percutaneous transluminal coronary angioplasty, coronary artery bypass graft or carotid endarterectomy. Episodes from patients with three or more of hyperlipidaemia, hypertension, diabetes mellitus or ≥ 80 years old were also considered to be high risk

^cIncidence rates were calculated in terms of per 1000 person-years. CIs were calculated using Byar's formula [23]

^dHRs and corresponding CIs were calculated using CMA Version 3 software

^eThe number of AMI, stroke and CV mortality cases do not sum to the number of MACE cases. For patients with more than one of the individual outcomes, the earliest individual outcome was counted towards the MACE total

incidences of AMI (HR 1.00, 95% CI 0.14–7.10) or stroke (HR 0.80, 95% CI 0.21–2.98) between the two groups, although estimates were imprecise and wide CIs were generated. The analyses from the Mini-Sentinel programme were subsequently replicated using two US administrative claims databases that did not contribute to Mini-Sentinel, IMS PharMetrics Plus and Truven (now IBM) MarketScan [33]. For both of these databases, no increased risks of AMI or stroke were found among users of mirabegron compared with users of oxybutynin (PharMetrics—AMI: HR 0.67, 95% CI 0.33–1.37; stroke: HR 0.62, 95% CI 0.34–1.13; MarketScan—AMI: HR 0.57, 95% CI 0.17–1.95; stroke: HR 0.69, 95% CI 0.30–1.62). The CV outcomes in these studies were not validated against source information; in contrast, our study adopted a more robust approach by validating potential adverse CV outcomes through direct linkage to registries or review by adjudication panels.

In the current study, the risks of CV mortality and all-cause mortality were no higher with mirabegron relative to

antimuscarinics. Although our study controlled for several risk factors for CV and mortality outcomes, it is possible that unmeasured risk factors for these specific events differed between users of mirabegron and antimuscarinics and these factors may explain the findings that suggest mirabegron may be associated with reduced mortality. Examples of such factors include frailty markers or social determinants of health, which are known predictors of mortality [34–36].

The current study has several strengths. Our investigation comprised a large, multinational source population, which to the best of our knowledge included the highest number of mirabegron users of any study conducted to date. In addition, our study investigated CV disease risk following the administration of drugs used for treating OAB in a variety of real-world clinical practice settings. The study also investigated outcomes that have been found to be reliably recorded in electronic healthcare databases. The presence of these outcomes was confirmed through direct linkage to registries, via medical record review or physician questionnaires.

Additionally, this study compiled a wide array of data on patient characteristics, health systems, drug utilisation and medical practices across populations in different countries, which enhances the generalisability of the study findings to the population of mirabegron users in routine clinical practice throughout Europe and the USA.

The current study also has limitations. Although PS methods were implemented to minimise the potential for confounding, residual confounding by unmeasured covariates cannot be ruled out. By updating the PS for each episode, time-dependent changes in key baseline covariates were used in the matching process to reduce potential confounding. In addition, it is possible that some residual confounding by lifestyle factors may have occurred as most of the data sources did not capture complete information on these risk factors. However, the secondary analyses that we conducted to assess the impact of potential confounding due to inadequately measured lifestyle factors showed minimal residual confounding of the observed effect estimates (< 9% bias). For patients contributing more than one treatment episode to the study, the analysis assumed independence of the treatment episodes. As 88% of patients contributed only one matched treatment episode, any impact of including additional episodes from the same patients would be relatively small. Patients were allowed to have different outcome types within the same episode or during subsequent episodes. As occurrences of AMI and stroke are known to affect the risk of subsequently developing other CV events [37, 38], we conducted a secondary analysis to censor all adverse CV outcomes upon occurrence of the first event. As stated above, this analysis yielded nearly identical results to the primary analysis and therefore it was assumed that the inter-related nature of the study outcomes had a minimal effect.

Additional limitations are related to the data sources and data capture on the real-world use of OAB medications. The current study included analyses from a variety of data sources, each set up for different purposes and over different time periods. For example, health outcomes that do not result in billed medical services were not identifiable from the US data. For the Danish, Swedish and CPRD data sources, data on drugs that were administered in hospitals or nursing homes, over-the-counter drugs or drugs provided as samples were not captured, although this may only constitute a small proportion of patients. Further, there was no information on the prescribed daily dosage or indication of use. We assume that the drugs were taken as prescribed, but they may have been used on an 'as-needed' basis. However, we believe these limitations are mitigated by the extensive

exposure data that were assessed in this study and validation of the study outcomes. Assessing the severity of hypertension or whether hypertension is controlled by medication and non-pharmaceutical treatments in studies like this one is challenging. This is because the data are obtained from existing data sources, where blood pressure results are generally not available and the proxies may not be accurate. Therefore, we did not conduct analyses to assess the effect of uncontrolled hypertension on the CV outcomes in this study. Furthermore, a European study that assessed the effectiveness of a direct healthcare communication letter to decrease the proportion of patients with uncontrolled hypertension prior to mirabegron initiation found that uncontrolled hypertension was uncommon before starting treatment with mirabegron even before the dissemination of the letter [39]. While it is possible that mirabegron might have influenced hypertension control in some patients, the previously mentioned results indicate that mirabegron initiation is quite rare in patients with uncontrolled hypertension, so there would be little to be learned from this type of analysis.

5 Conclusions

This PASS found no higher risk of MACE, AMI, stroke, CV mortality or all-cause mortality among current users of mirabegron compared with current users of antimuscarinics. Given the diverse nature of the study population, these study findings may be generalisable to mirabegron users in similar healthcare systems beyond those included in this study.

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Declarations

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Conflicts of Interest/Competing Interests Cheryl Enger and John D. Seeger are employees of Optum. Veena Hoffman and Kelesitse Phiri were employees of Optum at the time the study was conducted. Veena Hoffman, Kelesitse Phiri, Cheryl Enger and John D. Seeger hold stock in UnitedHealth Group, Optum's parent company. UnitedHealthcare, a UnitedHealth subsidiary, is a major purchaser of pharmaceutical products. The work was funded with a research contract between Optum and Astellas. Jesper Hallas and Morten Olesen have worked on a project commissioned by Astellas, with funding paid to their employer, the University of Southern Denmark. Nina Sahlertz Kristiansen was an employee of the University of Southern Denmark at the time the study was conducted. The contract granted the research team at the University of Southern Denmark independent publication rights. Marie Linder is an employee of the Centre for Pharmacoepidemiology, Karolinska Institutet. Ingvild Odsbu was an employee of the Centre for Pharmacoepidemiology at the time the study was conducted. The Centre receives grants from several entities (pharmaceutical companies, regulatory authorities and contract research organisations), including Astellas and Pfizer, for performance of drug safety and drug utilisation studies. Andrea V. Margulis, Alejandro Arana and Susana Perez-Gutthann are employees of RTI Health Solutions. RTI Health Solutions is a unit of RTI International, an independent non-profit organisation that conducts work for government, public and private organisations, including pharmaceutical companies. The RTI authors participated in this work in the course of employment as work for hire, pursuant to a contract to conduct an independent research study for a client (Astellas). The authors received no compensation other than their annual salary from their employer. Brandon T. Suehs and Yihua Xu are employees of Humana Healthcare Research, which received funding from Astellas in connection with the performance of this study. Libby Horter was an employee of Humana Healthcare Research at the time the study was conducted. Kwame Ap- penteng and Stefan de Vogel are employees of Astellas.

Ethics Approval The New England Institutional Review Board (IRB) approved the part of the study conducted by Optum. According to Danish law, register-based studies are exempt from approval by the Danish National Committee on Biomedical Research Ethics. Neither IRBs nor internal ethics committees are used/required at the University of Southern Denmark. The regional ethical board at Karolinska Institutet approved the part of the study conducted by Karolinska Institutet. The RTI International IRB determined that RTI researchers would not be engaged in research with human subjects. With this designation, additional review, approval or exemption by the RTI International IRB was not necessary. The Independent Scientific Advisory Committee for Medicines and Healthcare products Regulatory Agency database research approved the use of the CPRD database by RTI Health Solu-

tions. Advarra (formerly known as Schulman IRB) approved the part of the study conducted by Humana Healthcare Research.

Consent to Participate Not applicable.

Consent for Publication Not applicable.

Availability of Data and Material This observational study is based on individual patient data maintained by five separate data sources. Therefore, we are not able to make these data available to the public.

Code Availability Proprietary code was developed to perform the analyses on the data. The code applied to each data source and is maintained and archived according to the policies specific to each source. Researchers desiring access to the code would be required to obtain permission from the creator.

Authors' Contributions All authors provided substantial contributions to the conception or design of the study or the acquisition, analysis or interpretation of the data for the manuscript and the drafting of the manuscript or the revision of the manuscript for important intellectual content. The authors also provided final approval of the version to be submitted and agree to be accountable for all aspects of the manuscript.

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Authors and Affiliations

Veena Hoffman¹  · Jesper Hallas²  · Marie Linder³  · Andrea V. Margulis⁴  · Brandon T. Suehs⁵  · Alejandro Arana⁴  · Kelesitse Phiri¹  · Cheryl Enger¹  · Libby Horter⁵  · Ingvild Odsbu³  · Morten Olesen²  · Susana Perez-Gutthann⁴  · Yihua Xu⁵ · Nina Sahlertz Kristiansen²  · Kwame Appenteng⁶ · Stefan de Vogel⁷  · John D. Seeger¹  on behalf of the mirabegron PMR-PASS study group

¹ Present Address: Optum, 1325 Boylston Street, Boston, MA 02215, USA

² Present Address: University of Southern Denmark, Odense, Denmark

³ Present Address: Centre for Pharmacoepidemiology, Karolinska Institutet, Stockholm, Sweden

⁴ RTI Health Solutions, Barcelona, Spain

⁵ Present Address: Humana Healthcare Research, Louisville, KY, USA

⁶ Astellas Pharma US, Chicago, IL, USA

⁷ Astellas Pharma Europe B.V., Leiden, The Netherlands